Learning from Inventory Availability Information: 
Field Evidence from Amazon

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Many online retailers provide real-time inventory availability information. Customers can learn from the inventory level and update their beliefs about product quality. Thus, consumer purchasing behavior may be impacted by the availability information. Based on a unique setting from Amazon lightning deals, which displays the percentage of inventory consumed in real time, we explore whether and how consumers learn from inventory availability information. Identifying the effect of learning on consumer decisions has been a notoriously difficult empirical question due to endogeneity concerns. We address this issue by running a randomized field experiment on Amazon. In our experiment, we create exogenous shocks on the inventory availability information to a random subset of Amazon lightning deals. In addition, we track the dynamic purchasing behavior and inventory information for 23,665 lightning deals offered by Amazon in August 2016 and exploit their panel structure to further explore the underlying mechanisms behind learning. We find evidence of consumer learning from inventory information: a decrease in product availability causally attracts more sales in the future; in particular, a 10% increase in past sales leads to a 2.08% increase in cart add-ins in the next hour. Moreover, we show that buyers use observable product characteristics to moderate their inferences when learning from others, and customers primarily learn about the deal’s value rather than the product quality from the availability information.

Key words: Learning, inventory availability, field experiment, panel data, fixed effect, consumer behavior, Amazon, retail operations.

1. Introduction

Flash deal websites, such as Amazon Deals, Groupon or LivingSocial, have emerged as a popular means of selling products online. Using these platforms, sellers can advertise their products by offering them at a deep discount. New deals are announced on a daily basis, and they are only available over a specific period of time, ranging from a few hours to a few days. In the United States alone, consumers had spent approximately $7 million a day and more than $2.5 billion a year on flash deals by 2012 (Li and Wu 2014), and this number is estimated to exceed $3.9 billion
a year by 2015 (BIA/Kelsey 2011). Amazon, the largest U.S. online retailer, offers hundreds of limited-time deals every day, known as lightning deals. On July 15, 2015, Amazon promoted a 24-hour sale event called Prime Day, in which the peak orders of lightning deals surpassed those of Black Friday 2014. The popularity of flash deals has grown in many countries across the world, as well. In China, Alibaba’s Singles’ Day sales surged 60% to $14.3 billion on Nov 11, 2015.

Many flash deal websites provide inventory information in real time. It is important to understand consumer purchasing behavior during flash sales and the impact of this piece of real-time information. Inventory information signals previous customers’ preferences and opinions. By observing past purchasing decisions from inventory information, customers can draw inferences and update their beliefs about product quality, especially when their prior knowledge about the product is imperfect. This is called observational learning or herding. In a retail context, when deals advertise a product’s brisk prior sales or limited availability, observational learning causes an acceleration in subsequent sales. Consumers not only care about the presence of herding but also moderate their learning using various deal information, such as observable deal characteristics or the out-of-stock risk signaled by a product’s low inventory level.

In this paper, we explore whether consumers’ purchasing behavior is impacted by inventory availability information and which aspect of a sale customers primarily learn about from the inventory information when purchasing flash-sale products.

By showing a product’s sold inventory as a real-time percentage, Amazon’s lightning deals are an ideal research context in which to investigate whether consumers react to the inventory availability information. Lightning deals differ from traditional online sales in two distinctive ways. First, lightning deal sellers allocate a fixed amount of inventory to a lightning deal sale, and the deal page prominently displays the percentage of products sold in real time, which reveals product availability. By allowing customers to observe how much inventory has been consumed, this practice allows customers to reevaluate their desire for a product and could drive more sales for popular items. Second, unlike traditional online sales, lightning deals are sold at a discounted price for a limited time, e.g., 4 hours for most deals, and the remaining time is displayed as a countdown. Having to make decisions in a limited time under the pressure of missing the deal, customers need to rely on as much information about the product as possible. These features make inventory availability an important factor in their assessment; for example, New York Observer noticed that

“The countdown underneath each limited-time item on Amazon.com is both stressful and exciting. Better yet, seeing that 83 percent of the inventory has already been purchased is much more appealing.”


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We collect a proprietary and rich data set from Amazon.com by downloading all available information every 30 seconds. Besides the standard product information displayed on Amazon, such as customer reviews, price and promotional discount, we also collect the dynamic inventory information from a status bar under each deal indicating the percentage of that item currently in customers’ cart or purchased in real time. When a customer clicks the “Add to Cart” button, a claim is triggered, and the claimed status bar increases the relative percentage over the total predetermined inventory. We can, as a result, create a reliable metrics to measure cart-adding behavior given our fine downloading time resolution. We define the outcome variable as the potential sales, i.e., the cart add-ins. The independent variable is the inventory availability information, i.e., the percentage of deals claimed.

Our paper aims to estimate the causal relationship between product availability information and consumers’ purchasing behavior. Identifying the effect of learning on consumer decisions has been recognized as a notoriously difficult empirical question due to endogeneity concerns (Manski 1993 and Manski 2000). For instance, the observation that well-sold products attract more future sales may simply be due to the fact that products are heterogeneous and products with a high (low) quality are sold faster (slower). The main challenge is distinguishing learning from the unobserved heterogeneity across deals.

To address this issue, we employ a randomized field experiment on Amazon.com. In the experiment, we create an exogenous and instant shock in the inventory availability information to a random subset of Amazon lightning deals by adding deals to carts from multiple accounts. We conduct the field experiment over two weeks in September 2016. We then adopt a difference-in-difference approach to quantify the incremental cart add-ins across pre-treatment and post-treatment periods of the treated and controlled deals. Using our randomized treatments, we find the following causal relationship: a decrease in product availability, i.e., an increase in past sales, attracts more sales in the future. In particular, a 10% increase in past sales leads to a 2.08% increase in potential sales in the next hour. This result indicates that consumer learning has a significant and substantial impact on purchasing behavior.

To study the underlying mechanisms behind customer learning, we exploit the data set collected from Amazon.com in August 2016, which contains 23,665 lightning deals launched by Amazon in that month. We take advantage of the panel structure of our data and employ the fixed-effect model to capture any time-invariant, unobserved heterogeneity among different products. We find that buyers engage in rational learning: they attribute herding to external efforts rather than

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3 We also construct a proxy for sales by taking the difference between the hour-end number of claims and the hour-beginning number of claims. We find our results hold qualitatively across the sales measure. However, this measure is less accurate than cart add-ins and therefore is not included as the main outcome variable.
solely to intrinsic quality by using observable deal characteristics to moderate their inferences. Interestingly, we find that customers learn about the deal’s value rather than the deal’s quality from availability information: a deep discount rate weakens learning momentum because customers partially attribute herding to a good deal and might ignore the inference drawn from the inventory information; however, a good customer review does not weaken herding, which indicates that reviews and availability information signal different information content, and therefore, customers do not mainly learn about product quality from availability information. Further, we find that customers’ learning momentum increases when they face lower inventory availability (higher past sales), i.e., deals with low availability sell faster. This result is consistent with the scarcity effect—advertising a product’s low inventory level can create an out-of-stock pressure among customers, and thus prompt a resulting urgency to buy that product immediately.

Our study indicates that, in the online flash sales context, inventory availability is an important source of information that dictates consumer purchasing behavior. Past literature has documented how sales ranking or the absolute past sales volume affects future sales (Cai et al. 2009, Chen et al. 2011, and Li and Wu 2014). To the best of our knowledge, we are the first to establish the causal relationship between inventory information and consumer learning by running a field experiment in a retail context. Moreover, we complement the literature by shedding light on the underlying behavioral mechanisms behind consumer learning and pinpointing the specific aspects of a deal that a customer learns about from inventory information. More broadly, our results contribute to the literature on consumer learning and their strategic reaction to inventory information, which we summarize next.

2. Literature Review

Past studies have explored the impact of inventory information on consumer demand. Inventory information influences consumer purchasing decisions via both observational learning and scarcity. There is an extant literature focused on these effects.

Our paper is related to the stream of literature on observational learning and herding. The theoretical work shows that herding behavior—an individual may draw inferences from accumulative decisions by previous decision makers and follow their actions—can arise as an equilibrium (Banerjee 1992 and Bikhchandani et al. 1992). Stock and Balachander (2005) point out that the “scarcity strategy” adopted by companies signals a higher product quality and drives profits. Debo and Van Ryzin (2009) study the consumer herding behavior in a newsvendor context and propose an asymmetric inventory allocation strategy to increase satisfied demand. Veeraraghavan and Debo (2011) study a service context and find that herding can arise as a queue-joining equilibrium. Kremer and Debo (2015) use a lab experiment to show that uninformed customers infer product quality from wait time.
Our paper follows the recent interest in empirically quantifying herding in various contexts. In microloan markets, Herzenstein et al. (2011) and Zhang and Liu (2012) find evidence of herding among lenders, i.e., well-funded borrower listings tend to attract more funding. Using a natural experimental setting from an information policy shift at Amazon.com, Chen et al. (2011) focus on the digital cameras market and find that positive observational learning information increases sales, but negative observational learning information has no effect. In the context of restaurant dining, Cai et al. (2009) design a field experiment to distinguish the observational learning effect from the saliency effect. Using the data of daily deals on Groupon.com, Li and Wu (2014) explore the relationship between herding and social media word of mouth, and they find that the primary mechanism behind social media information is increasing product awareness. Our identification strategies are inspired by the above literature. To our knowledge, however, our paper is the first to study a causal effect of inventory availability information on consumer behavior by conducting a field experiment in the retail industry, and we are the first to distinguish which aspects of a deal customers learn about from inventory information.

Our work is also related to the stream of literature on scarcity. Scarcity can be a deliberate strategy for making a product more desirable. A low inventory level indicates that the product may not be available in the future, which can induce a “buying frenzy” behavior among customers (DeGraba 1995). Gallino et al. (2013) empirically explore the scarcity effect in the automotive industry. The authors show evidence of the scarcity effect when the added inventory does not increase the variety in model, but this effect vanishes as the variety increases. Recent research has proposed various strategies in response to consumers’ strategic reactions to scarcity. Liu and Van Ryzin (2008) show that sellers can create rationing risk by deliberately understocking products, and that the resulting threat of shortages induces customers to purchase earlier at higher prices. Cui and Shin (2016) show that sellers can adopt an aggregate inventory information-sharing strategy to reduce the shortage penalty cost by creating a stockout “illusion” among customers.

Further, our paper studies how consumers react to product availability information. In this sense, our study is closely related to the literature that studies consumers’ strategic reaction to product availability. The literature explores several decision-making mechanisms: (i) customers learn from the past stockout experience and anticipate a higher probability of a stockout on future orders, and (ii) the variety in inventory augments consumer choice options, which may confuse customers or enable a better preference match among heterogeneous customers.

In the face of a stockout, Netessine and Rudi (2003) consider a demand substitution behavior between competing sellers in a single-period model, and Netessine et al. (2006) analyze a backorder behavior in a multi-period game. Hall and Porteus (2000) and Gaur and Park (2007) study the change in retailers’ optimal inventory policy in response to strategic consumer learning when
customers react to the last stockout or remember the entire service history. Musalem et al. (2010) develop a structural demand model that estimates the effect of stockouts on consumer demand. Anderson et al. (2006) conduct a field experiment in a mail-order catalog to show that stockout adversely impacts both current demand as well as future demand. Craig et al. (2016) use a field experiment to show that an increased fill rate drives a significant increase in next-year sales. Tomlin (2009) investigates a retailer’s inventory and sourcing strategy when it learns about changes in the supplier’s service level. Kabra et al. (2015) study a bike-sharing context and empirically show that customers trade off bike availability and accessibility. Inventory variety has also been shown to be an important factor in dictating consumer buying decisions. Ryzin and Mahajan (1999), Gaur and Honhon (2006) and Cachon and Kök (2007) use consumer choice models to study the optimal assortment strategies, i.e., selecting a subset of variants to stock. Gallino et al. (2013) provide empirical evidence that product variety drives more sales for car dealers because customers are more likely to find an item that matches their preferences. On the other hand, Iyengar and Lepper (1999) and Iyengar and Lepper (2000) point out that high variety may create confusion in the decision-making process and thus lead to lower sales.

These papers offer innovative operations strategies in response to consumers’ past negative encounters or strategic decision making, whereas we focus on the consumer behavior aspect and provide causal empirical evidence of strategic learning in a retail context. Our paper shows that customers not only learn from real-time availability information, but also rationally use observable product attributes to moderate their inferences about the deal’s quality. This finding supports the literature’s theoretical modeling of rational learners and strategic decision making.

3. Background and Research Hypotheses

In this section, we demonstrate our research setting and theorize the research hypotheses for observational learning, and the behavioral mechanisms underlying consumer learning.

3.1. Research Background

Online flash sales provide a unique context to study how product availability influences consumers’ purchasing behavior. We choose the Amazon lightning deal platform, one of the largest daily deal sites, as our research setting.

Amazon launched “lightning deals” in 2009 — a marketing tool to raise awareness for a product or a brand. Amazon offers hundreds of lightning deals daily. These deals are deeply discounted (an average of 40%) and limited in time (lasting for 4 to 24 hours). A unique feature of lightning deals is that a seller allocates a fixed, limited amount of inventory to a lightning deal sale, and the page provides a real-time status bar indicating the percentage of available units that have already been claimed by customers, which reveals product availability. A timer appears below each lightning
Figure 1 Amazon Lightning Deal Example

$35.99
List: $69.99 (64% off)
10% Claimed Ends in 3:12:42
Small-Sized Thermo-Electric Dehumidifier
Sold by DIGGOT and Fulfilled by Amazon.

Add to Cart

$39.90
List: $49.90 (78% off)
29% Claimed Ends in 1:37:42
Woo Art Giclee Canvas Print Misty Lake Modern Canvas Wall Art for
Sold by WooArt and Fulfilled by Amazon.

Add to Cart

$35.99 - $47.99
46% Claimed Ends in 37:42
Elles Bedding Collection 500-
Thread Count Sateen Sheet Set Set...
Sold by Elles Bedding and Fulfilled by
Amazon.

Choose options

$9.99
List: $29.96 (67% off)
15% Claimed Ends in 1:57:42
Cutin8 Cheese Knife Set of 4,
Stainless Steel with Black Finish...
Sold by Innovative World and Fulfilled by
Amazon.

Add to Cart

deal, indicating how much time is remaining for the deal in real time. Like regular Amazon product pages, lightning deal pages also display standard product characteristics such as price and reviews. In Figure 1 we illustrate the information displayed on an Amazon’s lightning deal page.

3.2. Observational Learning

The literature on observational learning considers a situation where a market participant makes decisions based on two sources of information. One is her private knowledge of the product. This information is often imperfect, and thus, she is uncertain about the product value. The other is the information derived from the behavior of other market participants. When a subsequent decision maker observes predecessors’ decisions, she uses this information to update her belief about the product value. If the decision maker has limited private information, she may disregard her own information and simply follow others’ behavior. If the decision maker is very knowledgeable about the product, the reliance on others’ actions diminishes (Bikhchandani et al. 1992). Such behavior—using information acquired by watching others—is referred to as herding or observational learning.

Past observational learning literature has focused on a situation where all available actions are observable, e.g., each previous customer’s purchasing decision on a product. Recent studies have extended this assumption and shown that when agents observe aggregate prior decisions, e.g., the total number of customers who have purchased a product, herding on the aggregate observable choices can sustain as an equilibrium (Guarino et al. 2011).

In our research context, herding occurs when deals with higher prior sales tend to attract more sales in the future period. Intuitively, consider two products with the same characteristics; the deal
with higher prior sales sends a strong signal to subsequent customers, and an uninformed customer can make an inference of the deal quality and would expect the more popular deal to be more valuable than the alternative.

The two key drivers of herding are the uncertainty in buyers’ own information and their ability to learn from prior customers’ purchasing decisions. Amazon’s lightning deal platform provides an ideal research context that facilitates the mechanism of herding. First, more than 80% of deal buyers are new customers who are likely to be uninformed about the product and thus tend to herd when prior sales are high (Dholakia 2011). Second, by explicitly highlighting the percentage of products sold in real time, lightning deals provide a useful signal to customers which allows them to learn about the deal’s quality. Moreover, lightning deals run for only a limited number of hours. Having to make the buying decisions in limited time under the pressure of missing the deal, customers need to rely on as much information about the product as possible. These features make other customers’ decisions an important resource.

In short, our paper seeks to quantify the momentum of observational learning on the Amazon flash-sale platform. We expect that observational learning from inventory availability information significantly affects future sales.\(^4\)

### 3.3. Behavioral Mechanisms

The herding literature has shown evidence that observational learners care about not only the presence of herding, but also the various reasons that give rise to the herd (Zhang and Liu 2012). In other words, instead of passively replicating others’ purchasing decisions, consumers engage in rational learning: they attribute herding to external efforts rather than solely to intrinsic quality by using observable deal characteristics to moderate their inferences. Anderson and Holt (1997) and Goeree et al. (2007) show that decision makers can recognize a “mistaken” private signal that indicates the incorrect state, and that rational herding is formed in most periods in a lab experiment. Simonsohn and Ariely (2008) document the evidence of less rational herding among inexperienced eBay bidders when they herd into auctions with more existing bids but ignore the fact that the existing bids result from the no-longer-available lower starting prices rather than from higher quality. Rational learning has been documented in various practical contexts. For example, Duan et al. (2009) find that software buyers moderate their herding behavior on less popular products by factoring in the user rating. Zhang and Liu (2012) study lenders’ funding behavior in microloan markets and find evidence of rational herding: listing attributes such as credit scores

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\(^4\) When the product availability information is not shown to customers, there is no observational learning, and sales should appear constant over time after controlling for other factors. When the availability information is displayed, high prior sales signal the higher deal quality. Thus, the purchasing rate increases over time and the sales should appear as a convex, increasing shape over time.
signal the creditworthiness of a borrower, and rational lenders partially attribute the funding status to its attributes. Li and Wu (2014) identify social media information as another source of external effort that can moderate the herding momentum among deal buyers.

In our research context, rational learning occurs when consumer learning depends on the observable deal characteristics. Intuitively, consider two equally well-sold deals with identical observable characteristics except that deal 1 has a larger discount than deal 2. A rational subsequent buyer would think that the previous customers must possess positive private information if they chose deal 2 over deal 1. For example, previous customers might be experienced lightning deal buyers and know that deal 2 is more valuable because the product is rarely on sale. Therefore, a rational buyer would partially attribute deal 1’s sales to its deep discount, and draw a more positive incremental quality inference about the less-discounted deal.

More importantly, we investigate which aspects of a deal consumers learn about. The quality of a sale consists of two attributes: the product quality (i.e., whether the product rating is good enough) and the deal value (i.e., whether the discount is good enough). Amazon’s lightning deal setting allows us to gather empirical evidence on whether consumers mainly infer the deal value, the product quality or both from the inventory information.

If the deal value dictates a customer’s purchasing decisions, she uses the deal discount depth as a moderating factor in the inference obtained from availability information. Depending on whether customers learn about deal value from the availability information, the discount depth impacts customer learning in two opposing ways. If a customer learns about deal value from inventory information, then when seeing that a deeply discounted deal has a low inventory availability, she will partially attribute the good sales to the deal’s value (Zhang and Liu 2012). As a result, a deep discount weakens a deal’s herding momentum. On the other hand, if customers do not learn about the deal value from the inventory information, herding and the deal discount are two different information sources that provide different information content, and they may complement each other to generate sales (Kirmani and Rao 2000). In this situation, a higher discount amplifies a deal’s herding momentum.

Similarly, if product quality drives a customer’s purchasing decisions, she uses the product rating as an information source to moderate her inferences. There are two opposing hypotheses on how product rating impacts customer learning. By the same logic, one hypothesis suggests that when a customer learns about product quality from inventory information, she partially attributes large sales to a high product rating, and hence, a higher customer review weakens a deal’s herding momentum. The second hypothesis suggests that a higher rating can amplify a deal’s herding momentum when reviews and herding provide different information signals.
Consumer learning could also be driven by inventory scarcity. Scarcity affects sales through several mechanisms. DeGraba (1995) suggests that scarcity creates a buying frenzy among customers. Because low inventory availability indicates that the product may not be available in the future, consumers rush to purchase it under the out-of-stock pressure before they become informed. On the other hand, Stock and Balachander (2005) suggest that scarcity signals the popularity and quality of the product. Balachander et al. (2009) find empirical support for the signaling theory in the automobile market. Brock (1968), Lynn (1991) and Tereyagoglu and Veeraraghavan (2012) also point out that customers may prefer a product that is more exclusive, and thus, scarcity is associated with a stronger preference.

In our context, when product availability is relatively high, consumer learning—i.e., higher past sales drive more future sales—is primarily driven by herding. When product availability is relatively low, the scarcity effect begins to kick in and consumer learning is driven by both herding and scarcity. Therefore, in the presence of the scarcity effect, we expect that the average learning momentum increases as customers respond to an increasing average percentage of claimed products.

In short, we expect customers to rely on observable deal characteristics to moderate inferences, and we further investigate customers’ strategic learning of two important aspects of the deal quality—deal value and product quality. In addition, we study the impact of a low inventory level—the scarcity effect—on purchasing behavior.

4. Data

We collected a proprietary data set from Amazon.com by downloading all available information from Amazon’s lightning deal platform every 30 seconds in August and September of 2016. In addition, we kept downloading the data while we ran a field experiment in September 2016. Despite Amazon’s being one of the largest flash sales sites in the world, our paper is (to our knowledge) the first one to gather and take advantage of this data source. We gather two portions of data from Amazon lightning deals: (i) real-time information and (ii) static product characteristics.

Real-time inventory information. When a customer clicks the “Add to Cart” button, a claim is triggered, and the claimed status bar increases the relative percentage over the total predetermined inventory. If the customer does not complete the purchase within 15 minutes, Amazon will drop the product from the cart and drop the corresponding percentage of units claimed. Since each customer is restricted to purchasing only one unit from each lightning deal, the percentage increase reflects the number of customers who intend to purchase the deal. Figure 2 in the Appendix shows an example of the dynamic change of the percentage claim information over time for a lightning deal.

Our granular download-time resolution (i.e., 30 seconds) allows us to capture the minimum increment (or decrement) of product availability. We use it to calculate the number of customers
who add the product to their carts within a period of time. We can also use it to impute the approximate inventory allocated to a deal.\(^5\) In addition, we collect the remaining time from the timer displayed under each lightning deal.

For each deal, we construct the outcome variable *Cart add-ins* as the total positive increment in the percentage claimed in a specific time period.\(^6\) When a customer adds an item to the cart, she expresses her interest in the product and creates a chance to make a purchase. Cart add-ins, as a result, measure potential customers who have expressed interest in purchasing.

**Product characteristics data.** For each deal, we also gather static product information. Specifically, we collect the final discounted price, the listed original price, the actual original price, the product rating, the number of reviews and the number of options (in color or size). Lightning deals display the listed original price and a corresponding discount percentage below each deal. However, the listed original price is the manufacturer’s suggested retail price, which is often higher than the non-promotional price on Amazon. Therefore, we also collect the actual original price from Amazon, from which we calculate the actual discount rate. We obtain two measures of the promotional depth: the actual discount and the listed discount. These allow us to investigate whether customers are knowledgeable enough to differentiate them and which discount rates dictate decision making.

**Descriptive statistics.** Table 1 presents the summary statistics for the dynamic variables and the time-invariant product characteristics in August 2016. Our sample includes 23,665 deals that run for an average of 5.2 hours. The average cart add-ins per hour are 4.47%.\(^7\) The average final deal price is $24.15 with an actual original price of $31.67 (discount percentage of 22%) and a listed original price of $43.73 (discount percentage of 39%). On average, a deal has 130.37 reviews and an average rating of 3.91 out of 5 stars. The sellers allocate more than 35.47 units of inventory to lightning deal sales (given we underestimate the inventory level beyond 100 units).

5. **Identification Strategies**

We aim to estimate the causal relationship between product availability and consumers’ future purchasing behavior. Identifying the effect of social learning on consumer decisions has been recognized as a notoriously difficult empirical question due to endogeneity concerns (Cai et al. 2009). The main problem is in distinguishing social learning from unobserved heterogeneity across deals. A common unobserved heterogeneity is the deal quality. Deals with good quality tend to have high

\(^5\) We calculate inventory as \(100/(\text{the minimum increment})\). For example, if the minimum increment is 2%, the inventory level is 50 units. Please note that the minimum increment takes only integers. Therefore, we approximate the inventory level as 100 for deals with more than 100 units.

\(^6\) Given our granular download-time resolution and one-customer-one-deal purchasing restriction, *Cart add-ins* is an accurate and good measure of potential customer demand.

\(^7\) The max cart add-in is 228%. It indicates that for this product in a specific hour, many customers have shown interest in purchasing but eventually dropped the item.
past sales and high future sales even without learning, while some other deals are less popular, and in the most extreme case, may have zero sales across the selling period. In this case, the ordinary least squares (OLS) regression may attribute high future sales to high past sales when it could actually be driven by deal quality. Therefore, the OLS estimator may overestimate the social learning effect and suffer from the omitted variable bias.

Following the past empirical literature, we address this issue by using two identification strategies: (i) randomized field experiment and (ii) panel data analysis. The randomized field experiment randomly assigns the treatment to products or individuals; Cai et al. (2009) and Chen et al. (2011) have adopted randomized experiments to estimate the social learning effect in other research contexts. The advantage of using randomized experiments is that, since the treatment is unconditionally randomly assigned, the estimator represents an unbiased causal effect. In other words, the estimators obtained from randomized control experiments have high internal validity (Levitt and List 2007). However, because field experiments are often conducted under financial and human resource constraints, the number of observations is often relatively small, which limits the external validity.

The panel data approach exploits the panel structure of the data and uses individual-level fixed effects to control for time-invariant unobserved heterogeneity; Sorensen (2006), Duan et al. (2009), Zhang and Liu (2012) and Li and Wu (2014) have adopted the fixed effect specification. The advantage of using this approach is that the panel data set is often large and covers various products or individuals. Thus, the estimators derived from the panel approach have high external validity. However, because the causal identification obtained by the fixed-effect model requires the assumption that the unobserved heterogeneity is time-invariant, the fixed-effect specification often lacks the internal validity and needs various robustness checks to prove this assumption.

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<td>Duration</td>
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<td>0.98</td>
<td>3.00</td>
<td>6.00</td>
<td>23,665</td>
</tr>
<tr>
<td></td>
<td>Inventory</td>
<td>35.47</td>
<td>40.12</td>
<td>1.00</td>
<td>100.00</td>
<td>23,665</td>
</tr>
</tbody>
</table>
Our paper employs both identification strategies to establish both the internal and external validity of our findings. In particular, we first conduct a randomized field experiment to quantify the causal impact of consumer learning on sales among a set of Amazon deals. We then turn to the panel data set to re-estimate this causal effect over a much larger set of deals. By demonstrating that these two estimators are similar in directions and magnitudes, we not only provide a crucial understanding of consumer learning but we also show that our findings are equally descriptive of the world at large.

Given the average inventory level for lightning deals is relatively high (e.g., the average maximum fraction claimed is 10.74%), consumers do not face a lot of stockouts, and the primary driver in consumer learning is herding. The scarcity effect begins to come into play when availability becomes relatively low.\(^8\)

**Randomized field experiment.** We run a field experiment on Amazon. In our experiment, we generate a shock in the claim information by creating an instance spike in sales—which is our treatment—to a randomly selected group of Amazon deals. This spike lasts for one hour. We started our experiment on average two hours before the deal period ended. This allows us to compare the purchasing behavior across three periods: pre-treatment, treatment and post-treatment periods. We adopt a difference-in-difference analysis to quantify the incremental cart-adding behavior driven by our treatment. Because the deals receive a treatment two hours before their ending time, the average inventory availability is far from stockout and it is unlikely that customers face an out-of-stock pressure. Therefore, the field experiment identifies the learning momentum mainly driven by herding.

**Panel data approach.** Our proprietary data set has a panel structure, where for each deal, the percentage claimed varies dynamically over time. The panel structure allows us to use a fixed-effect model to capture any time-invariant unobserved heterogeneity; in this way, we exploit the variation within each deal to obtain an estimate of learning and its interaction with product attributes. Using this fixed-effect approach, we show the external validity of our findings in the field experiment given that the panel data covers a much broader spectrum of products. Further, we interact the availability information with observable product characteristics to explore possible mechanisms behind learning. Because the fixed-effect model estimates the average learning momentum across various inventory availability levels, the learning momentum may include both herding and scarcity effects.

\(^8\) Note that we cannot completely disentangle the herding effect from the scarcity effect as it is not well understood when the scarcity effect begins to kick in (e.g., the availability level below which customers start to care about scarcity).
6. Randomized Field Experiment

In this section, to identify the causal effect of the availability information on purchasing behavior, we design and conduct a randomized field experiment on Amazon’s lightning deal platform.

6.1. Experiment Design and Data Summary

Recall that an increment in claim is triggered by clicking the “Add to Cart” button. We take advantage of this mechanism and create an exogenous shock to the availability information. To this end, we create 10 Amazon accounts and add products to these accounts’ carts in a short period of time (i.e., within 3 minutes) to spike the claim information. Before the start of the experiment, we download all the available deals with a duration of at least two hours. There are on average 70 such deals available per hour. We randomly assigned an average of 40% of the deals to the treatment group that receives a shock in “past sales” and the rest of the deals to the control group. Figure 3 in the Appendix illustrates our field experiment design.

With the help of four research assistants, we were able to add the treated deals to the cart of each of the 10 accounts within 3 minutes. Figure 4 in the Appendix illustrates how the availability information changes when we spike cart add-ins in the experiment. Because Amazon drops products out of the cart after 15 minutes if the purchase has not been completed, we re-added the treated deals every 15 minutes. We started the experiment at either 2 p.m. or 3 p.m. in the afternoon and kept the shocks on each deal for around 60 minutes on one day (the 60 minutes include 45 minutes effective in-cart time and around 15 minutes operating time). We repeated the experiment six times: on Tuesday, Friday and Saturday of two consecutive weeks.9

There are in total 181 deals in the treatment group and 264 deals in the control group. Table 2 summarizes the deal characteristics, inventory information dynamics across the treatment group and the control group within a 50-minute window before the experiment, a 60-minute window within the experiment and a 50-minute window after the experiment. On average, we created an exogenous average increment of 16.48% in the percentage claimed to the 181 treated deals at the beginning of the treatment period. Table 2 also demonstrates that the deals in the treatment and control groups have similar characteristics such as hours to the end, listing discount percentage, actual discount percentage, review ratings and number of reviews, which proves that the treatment assignment is randomized. Table 2 provides initial evidence of consumer learning: in the treatment period, the treated deals receive a higher number of cart add-ins than the control deals.

9 An experiment spanning Tuesday, Friday and Saturday captures the effects on both weekdays and weekends.
### Table 2 Field Experiment Summary Statistics

<table>
<thead>
<tr>
<th>Window</th>
<th>Variable</th>
<th>Treatment (N=181)</th>
<th>Control (N=264)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Mean</td>
<td>Std. Dev</td>
<td>Mean</td>
</tr>
<tr>
<td>Pre-50-min</td>
<td>Start claim</td>
<td>5.18</td>
<td>8.29</td>
</tr>
<tr>
<td></td>
<td>Cart add-ins</td>
<td>7.13</td>
<td>9.40</td>
</tr>
<tr>
<td></td>
<td>Hour to end</td>
<td>3.44</td>
<td>0.50</td>
</tr>
<tr>
<td></td>
<td>List discount pct</td>
<td>0.62</td>
<td>0.20</td>
</tr>
<tr>
<td></td>
<td>Actual discount pct</td>
<td>0.23</td>
<td>0.11</td>
</tr>
<tr>
<td></td>
<td>Review rating</td>
<td>4.08</td>
<td>1.35</td>
</tr>
<tr>
<td></td>
<td>No. of reviews</td>
<td>97.88</td>
<td>255.82</td>
</tr>
<tr>
<td>Experiment</td>
<td>Start claim (before shocks)</td>
<td>9.14</td>
<td>13.04</td>
</tr>
<tr>
<td></td>
<td>Start claim (after shocks)</td>
<td>25.62</td>
<td>19.20</td>
</tr>
<tr>
<td></td>
<td>Cart add-ins</td>
<td>10.49</td>
<td>9.88</td>
</tr>
<tr>
<td>Post-50-min</td>
<td>Start claim</td>
<td>19.42</td>
<td>21.09</td>
</tr>
<tr>
<td></td>
<td>Cart add-ins</td>
<td>5.88</td>
<td>7.85</td>
</tr>
</tbody>
</table>

### 6.2. Experiment Results

We analyze the treatment effect of availability information on consumer cart-adding behavior. We first employ a pre-post comparison specification. Specifically, for each deal, we compare the total cart add-ins within the experiment time window with those within 50 minutes before (or after) the experiment. The validity of this specification relies on the assumption that the potential demand and consumer purchasing behavior are time-invariant. To control for potential heterogeneity over time, we then employ a difference-in-difference approach.

#### 6.2.1. Pre-post treatment comparison

Let $Treated\ Deal_i$ be a dummy variable to denote whether deal $i$ was assigned in the treatment group; $Treated\ Deal_i = 1$ represents the deals that received the treatment on inventory availability. Let $Treated\ Hour_p$ denote the treated hour dummy, where $p \in \{\text{pre-50-min period, experiment period, post-50-min period}\}$; $Treated\ Hour_p = 1$ represents period $p$, the experiment hour. We estimate the treatment effect by the following specification,

$$s_{i,p} = c + \alpha Treated\ Hour_p + T_i + Z_i + e_{i,p}, \ \forall i \ s.t. \ Treated\ Deal_i = 1 \ (1)$$

where $s_{i,p}$ denotes the number of times deal $i$ was added to the cart per minute within period $p$, i.e., cart-adding rate.\(^\text{10}\) In addition, $T_i$ control for time characteristics, e.g., the day of the week and the hour of the day effects, and $Z_i$ control for deal characteristics, e.g., review rating, total number of reviews, and discount depth. Please note that, in the randomized experiment, we do not need to

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\(^{10}\) The reason we measure the outcome variable by the cart-adding rate is that we vary the duration of the pre-experiment control period and the post-experiment control period as a robustness test. In order to make an accurate and correct comparison, we adopt the cart add-ins per minute as the dependent variable.
add any control covariates to obtain an unbiased estimate of the treatment effect because treatment is unconditionally randomly assigned; the addition of controls, however, can make the estimates more efficient (Paulsen and Smart 2013). Our main analysis will not include deal characteristics, but our analysis is qualitatively and quantitatively robust after including the deal characteristics in Table 6 in the Appendix.

The consumer learning effect can be measured by the incremental potential sales driven by the treatment, i.e., the coefficient of Treated Hour\_p. The validity of this specification resides in the assumption that the incoming potential demand and consumer purchasing behavior do not systematically change over time (conditional on our control variables),

Assumption 1. Pre-post comparison assumption: \( \mathbb{E}(e_{i,p} \mid \text{Treated Hour}_p) = 0. \)

We let \( t = \{\text{pre-50-min period, experiment period}\} \) in specification (1) to compare the treated period with the pre-50-min window. At the end of our experiment period, the fictional add-ins from our treatment are dropped from the carts, and as a result, the treatment effect diminishes in the 50-min period after the experiment. To show the robustness of our treatment effect over the period after the experiment, we re-run specification (1) where \( t = \{\text{experiment period, post-50-min period}\} \).

The estimation results are displayed in Columns I and IV in Table 3. As a placebo test, we repeat the above pre-post comparison analysis for the deals in the control group, i.e., for \( i \) that Treated Deal \(_i = 0\), to detect any systematic differentiation of the cart-adding rate over time. Columns II and V in Table 3 present the estimation results.

Column I shows evidence of consumer learning: a product’s brisk prior sales causally drive an acceleration in sales in the future period. Column II shows that for deals in the control group, the cart-adding rate within the 60-minute experiment window is not significantly different than during the 50-minute window prior to the experiment. In other words, Column II demonstrates that Assumption 1 is satisfied between pre-experiment and experiment periods, which confirms that the treatment effect identified in Column I is driven by our treatment rather than other factors. The main finding holds in Column IV, i.e., the treatment effect is statistically significant on the cart-adding rate over the post-experiment period. However, this treatment effect may be confounded by temporal heterogeneity in cart-adding behavior since Column V shows that the post-treatment period has a systematically greater cart-adding behavior compared to the experiment period, even for deals in the control group. Therefore, to control for such time heterogeneity, we next employ a difference-in-difference analysis.

6.2.2. Difference-in-difference (DiD) analysis. The difference-in-difference estimator is derived from the following specification,

\[
s_{i,p} = c + \text{Treated Deal}_i + \text{Treated Hour}_p + \beta \text{Treated Deal}_i \times \text{Treated Hour}_p + T_i + Z_i + e_{i,p}, \quad (2)
\]
Table 3  Field Experiment Analysis

<table>
<thead>
<tr>
<th>Specification</th>
<th>Pre-50-min Comparison</th>
<th>Post-50-min Comparison</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>I. Treated</td>
<td>II. Control</td>
</tr>
<tr>
<td>Treated Period</td>
<td>0.045**</td>
<td>−0.012</td>
</tr>
<tr>
<td></td>
<td>(0.017)</td>
<td>(0.018)</td>
</tr>
<tr>
<td>Treated Deal</td>
<td>−0.013</td>
<td>(0.019)</td>
</tr>
<tr>
<td>Treated Hour ×</td>
<td>0.057**</td>
<td>(0.026)</td>
</tr>
<tr>
<td>Treated Deal</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Saturday</td>
<td>0.008</td>
<td>0.080**</td>
</tr>
<tr>
<td></td>
<td>(0.030)</td>
<td>(0.035)</td>
</tr>
<tr>
<td>Tuesday</td>
<td>0.034</td>
<td>0.010</td>
</tr>
<tr>
<td></td>
<td>(0.029)</td>
<td>(0.033)</td>
</tr>
<tr>
<td>Hour</td>
<td>−0.020**</td>
<td>0.022***</td>
</tr>
<tr>
<td></td>
<td>(0.008)</td>
<td>(0.008)</td>
</tr>
<tr>
<td>Constant</td>
<td>0.341***</td>
<td>−0.145</td>
</tr>
<tr>
<td></td>
<td>(0.100)</td>
<td>(0.106)</td>
</tr>
<tr>
<td>Observations</td>
<td>362</td>
<td>528</td>
</tr>
</tbody>
</table>

Note: This table reports the estimated coefficient values of equations (1) and (2) with the time controls. Column I reports the pre-post comparison relative to the pre-50-min period for the treated deals of equation (1); Column II reports the pre-post comparison relative to the pre-50-min period for the control group of equation (1); Column IV reports the pre-post comparison relative to the post-50-min period for the treated deals of equation (1); Column V reports the pre-post comparison relative to the post-50-min period for the control group of equation (1); Columns III and VI report the estimation results from the difference-in-difference approach in equation (2). The errors are corrected for heteroskedasticity. Significance at *p<0.1; **p<0.05; ***p<0.01.

where the learning effect can be measured by the incremental cart-adding behavior driven by the treatment after removing the time effect, i.e., the coefficient of Treated Deal_i × Treated Hour_p. The key identification assumption for any DiD strategy is the parallel trends assumption: the outcome in treatment and control groups would follow the same time trend in the absence of the treatment. In other words, the unobserved characteristics of deals in the treatment and control groups are independent of the treatment, which we summarize below. Because our treatment is unconditionally randomly assigned (i.e., E(e_i,p | Treated Deal_i) = 0 ∀i,p), the parallel trend assumption is satisfied by the experiment design.

Assumption 2. DiD assumption: E(e_i,p | Treated Hour_p × Treated Deal_i) = 0 ∀i,p.

In our experiment, the treatment, i.e., the shocks in the percentage claim information, is not very likely to create a cannibalization effect and impact purchasing behavior for the control deals. This is because Amazon offers a wide product variety, and the types of products sold around the same time are quite different.
Columns III and VI estimate the treatment effect using the difference-in-difference specification in Equation (2) relative to the pre-50-minutes time period and the post-50-minutes time period, respectively. After removing the potential time heterogeneity, our DiD estimates remain significant, which suggests that a spike in past sales causally increases future sales. This provides empirical evidence for the existence of consumer learning.

Recall that we created an exogenous increment of 16.48% in the inventory claim bar. The result of the difference-in-difference specification in Columns III and VI indicates that such an increment translates into an increase of 3.42% cart add-ins in an hour (0.057% units per minute × 60 minutes). In other words, a 10% increase in the inventory claim bar leads to a 2.08% increase in cart add-ins. This provides empirical evidence that customers react to and learn from the inventory information, and the impact is significant and sizable. As a robustness test, we re-run our results using different time windows, i.e., 5 minutes, 10 minutes, 20 minutes before and after the experiment. The treatment effect is also robust to these time-window specifications. In Section 7, we will compare this estimate with the estimates obtained from the panel data analysis and examine whether estimates based on these two approaches are consistent in magnitude.

7. Panel Data Analysis

In this section, we exploit the panel structure of the data to estimate the learning momentum and the underlying behavioral mechanism during the purchase of flash-sale products.

7.1. Fixed-Effect Specification

In our collected data set, for each deal, the inventory information varies dynamically over time, but the product characteristics are time-invariant, which provides us with a panel structure of the data. The panel structure allows us to exploit the fixed-effect specification.

We denote \( t \) as the hour from the start of the deal, where \( t = 1, \ldots, T \); \( T \) is the duration of a deal. We use \( C_{i,t} \) to denote the total percentage of inventory claimed for deal \( i \) at the beginning of hour \( t \), and we use \( y_{i,t} \) to denote the total cart add-ins during hour \( t \), i.e., potential sales. The total amount of inventory claimed reflects the previous customers’ collective opinions regarding the worthiness of the deal and the product’s quality. The learning effect can be measured by the dependency of future sales on current inventory information, i.e., the coefficient of \( C_{i,t} \) on \( y_{i,t} \).

We first conduct a preliminary OLS analysis by looking at the correlation between future cart add-ins and current inventory information. The analysis tests whether \( y_{i,t} \) is positively correlated with the inventory information \( C_{i,t} \) after controlling for time-varying attributes \( X_t \) and time-invariant attributes \( Z_i \),

\[
y_{i,t} = \gamma C_{i,t} + \beta_1 X_t + \beta_2 Z_i + e_{i,t}. \tag{3}
\]
The time-varying attributes $X_t$ include the hour-of-day and day-of-the-week fixed effects. These attributes capture the possibility that sales tend to concentrate in certain hours of a deal or on certain days of the week. Time-invariant deal attributes $Z_i$ include the No. of Reviews, the number of reviews under a deal during the period when the deal is active; Review Rating, the average star rating of the reviews; Discount, the actual and listed promotional depth of the deal; Variety, a dummy variable equal to one if the product has multiple options, e.g., colors or sizes. The coefficient of $C_{i,t}$ tests the correlation between $C_{i,t}$ and $y_{i,t}$.

The available data may not capture all the heterogeneity across the deals. For example, a deal with a detailed product description or professional-looking photos is likely to attract more customers, but our data does not include these variables. Statistically, there may be unobserved heterogeneity in the error term, i.e., $e_{i,t} = u_i + \epsilon_{i,t}$, where $\epsilon_{i,t}$ is orthogonal of all the independent variables and $u_i$ represents the unobserved deal attributes such as the inherent quality of the deal. Since $u_i$ is correlated with other deal-specific features, $C_{i,t}$ and $Z_i$, without controlling for $u_i$, the OLS regression in equation (3) suffers from the omitted variable bias (Wooldridge 2010). Fortunately, the panel structure of the data allows us to use a fixed-effect specification to capture any time-invariant unobserved deal heterogeneity $u_i$,

$$y_{i,t} = \gamma C_{i,t} + \beta_1 X_t + \beta_2 Z_i + u_i + \epsilon_{i,t}. \quad (4)$$

The key identification assumption for the fixed-effect model is that unobservable deal heterogeneity is time-invariant, which we summarize below.

**Assumption 3.** Fixed-effect assumption: $\mathbb{E}(\epsilon_{i,t} | C_{i,t}, u_i) = 0 \forall i, t$.

This is plausible in our research context because our study examines the learning effect on sales over the course of a 4- to 6-hour deal. It is unlikely that the characteristics of the deal would change during the time period when the deal is active. Therefore, by controlling for the unobserved heterogeneity $u_i$, we identify consumer learning using within-deal variations in $C_{i,t}$, $y_{i,t}$ and $X_t$.

Moreover, consumers rely on observable deal characteristics to moderate their inferences from observing the inventory information. To explore the underlying behavioral mechanism of learning, we examine the cross-sectional variations in the observable listing attributes, which allow us to distinguish whether consumers primarily learn of the product quality or the deal quality. We augment equation (4) by including the interaction term between the cumulative inventory information and the observable deal characteristics like customer reviews and deal discount.

---

11 We downloaded Amazon’s data every 30 seconds, which allows us to examine the variation of reviews over the 4 hours when the deal is active. We found that the review properties do not change much during the lightning deal period. Therefore, the number of reviews and product rating are time-invariant variables.
\[ y_{i,t} = \gamma C_{i,t} + \beta_1 X_t + \beta_2 Z_i + \beta_3 C_{i,t} \times Z_i + u_i + \epsilon_{i,t}. \]  

(5)

The deal’s value, i.e., the discount level, impacts customer learning in two opposing ways. When a customer sees that a minimally discounted deal has a low inventory availability, she will argue that there must be something good about this deal that attracts other customers despite its small discount, and make a more positive incremental quality inference about these minimally discounted deals from the inventory information. Therefore, the learning momentum can be accentuated by a low discount rate, i.e., the coefficient of \( C_{i,t} \times \text{Deal discount} \) should be negative. In contrast, the opposite relation would suggest the discount rate and availability information provide different information content, as they complement each other to attract sales, and thus, it suggests that customers do not mainly learn about deal value.

Product quality, which is reflected in customer reviews, may also drive consumers’ learning in two opposing ways. When a customer sees that a low-rated deal has a low inventory availability, she will justify the herd with some other positive characteristics of the deal, and make a more positive incremental quality inference about it. Therefore, the learning momentum can be accentuated by a low product rating, i.e., the coefficient of \( C_{i,t} \times \text{Review rating} \) should be negative. Similarly, the opposite sign would suggest flash-sale buyers do not mainly learn about product quality from the availability information.

7.2. Estimation Results

Table 4 provides the estimated coefficient with potential sales as the dependent variable. Column I displays the results for the OLS model in equation (3). Column II displays the results for the fixed-effect model in equation (4). Columns III - VI report the results for the interacted fixed-effect model in Equation (5) with respect to customer reviews, deal discount rate and the option variety.

The OLS estimator in Column I suggests that the effect of inventory information \( (C_{i,t}) \) is positive and significant on sales in the next hour. However, the OLS estimates may suffer from the omitted variable bias and overestimate the learning effect. Below we report the estimates from the fixed-effect model.

7.2.1. Consumer learning. Column II of Table 4 reports the coefficient of the past inventory information for the fixed-effect model with time controls. As expected, the OLS specification in Column I overestimates the effect compared with the fixed-effect specification in Column II. After controlling for deal-specific heterogeneity, \( \text{Claim} \) has a significant and positive coefficient, which shows the existence of consumer learning—i.e., the availability information indeed influences following customers’ choices and a deal with higher past sales attracts more purchases in the future. More importantly, the magnitude of the causal effect identified by the panel data analysis aligns
### Table 4  Fixed-effect Analysis

<table>
<thead>
<tr>
<th>Specification</th>
<th>I</th>
<th>II</th>
<th>III</th>
<th>IV</th>
<th>V</th>
<th>VI</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Claim</strong></td>
<td>0.452***</td>
<td>0.299***</td>
<td>0.156***</td>
<td>0.220***</td>
<td>0.243***</td>
<td>0.215***</td>
</tr>
<tr>
<td></td>
<td>(0.002)</td>
<td>(0.023)</td>
<td>(0.043)</td>
<td>(0.044)</td>
<td>(0.045)</td>
<td>(0.044)</td>
</tr>
<tr>
<td><strong>No. of reviews</strong></td>
<td>−0.0001**</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.0001)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Review rating</strong></td>
<td>0.252***</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.017)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Variety</strong></td>
<td>−2.272***</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.060)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Actual discount</strong></td>
<td>0.005***</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.001)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Claim × No. of reviews</strong></td>
<td></td>
<td>0.00001</td>
<td></td>
<td>0.00001</td>
<td></td>
<td>0.00001</td>
</tr>
<tr>
<td></td>
<td></td>
<td>(0.00001)</td>
<td></td>
<td>(0.00001)</td>
<td></td>
<td>(0.00001)</td>
</tr>
<tr>
<td><strong>Claim × Review rating</strong></td>
<td>0.032***</td>
<td></td>
<td>0.029***</td>
<td></td>
<td>0.032***</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.008)</td>
<td></td>
<td>(0.009)</td>
<td></td>
<td>(0.008)</td>
<td></td>
</tr>
<tr>
<td><strong>Claim × Variety</strong></td>
<td>−0.111***</td>
<td>−0.114***</td>
<td>−0.111***</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.013)</td>
<td>(0.013)</td>
<td>(0.013)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Claim × Actual discount</strong></td>
<td>−0.002***</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.006)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Claim × List discount</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.0003</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>(0.0002)</td>
</tr>
</tbody>
</table>

| **Time control**              | Yes   | Yes   | Yes   | Yes   | Yes   | Yes   |
| Observations                  | 123,115 | 123,115 | 123,115 | 123,115 | 123,115 | 123,115 |

Note: This table reports the estimated coefficient values of equations (3), (4) and (5) and standard errors clustered at the deal level with the hour of the deal controls. Column I reports the sequential correlation results of the OLS regression of equation (3); Column II reports the results of the fixed-effect regression of equation (4); Column III reports the results of the fixed-effect model interacted with number of reviews and review rate; Column IV reports the results of the fixed-effect model interacted with actual discount; Column V reports the results of the fixed-effect model interacted with number of review, review rate and actual discount; finally, Column VI reports the results of the fixed-effect model interacted with all deal characteristics. The errors are corrected for heteroskedasticity.

Significance at *p<0.1; **p<0.05; ***p<0.01.

with the magnitude identified by the randomized field experiment. The panel data analysis suggests that, all else equal, a 10% increase in the cumulative percentage claimed leads to a 2.99% increase in the potential sales in the next hour. Recall that the field experiment suggests a 2.08% increase in the next hour’s cart add-ins. These two estimates are highly consistent in both direction and magnitude, which helps establish internal validity for the fixed effect approach by strengthening the causal argument and helps establish external validity for the field experiment by extrapolating the finding over a larger data set.

### 7.2.2. Behavioral mechanisms.

We find that consumer learning depends on the deal attributes, suggesting that consumers not only learn from other customers but also moderate their
inferences using various observable deal characteristics that give rise to the herd. The sign and the magnitude of the interaction terms are robust across Columns III - VI.

We show that a higher actual discount rate mitigates the learning effect, i.e., the actual discount has a negative interaction with Claim. In other words, the same herding momentum signals a better deal value if the deal has a weaker promotional depth. This finding suggests that discount rate and percentage-claimed information provide similar information content, and thus, customers learn about the deal value from inventory information.

Interestingly, the listed discount by deal sellers has no significant impact on learning, i.e., the interaction term of Claim × List discount in Column VI is insignificant. Recall that lightning deals’ listed discount percentage is almost twice as much as the actual discount percentage. This finding provides side evidence that customers are rational decision makers—they search for the actual regular price or the price offered by other sellers before making a purchase and they are knowledgeable enough to differentiate the actual discount rate from the listed discount rate.

We also show that a higher product rating amplifies the learning effect, i.e., the rating has a positive interaction with Claim. This suggests that customer rating and the availability information signal different information content, as these two sources of information complement each other to attract sales. Thus, customers do not learn about product quality from inventory information. In addition, the number of reviews has no significant effect on potential sales. This indicates that compared with the sentiment toward the products, the number of reviews does not have a significant influence on consumers’ choices.

Finally, the interaction term Claim × Variety is negative, which suggests that the signal strength of the availability information diminishes when customers choose from a larger number of options. Amazon’s lightning deal page displays the average inventory information across multiple options of deal. Customers have to click the “choose option” button and select the exact option they want. This costly action may dilute customer attention to the inventory information and thus weaken the strength of the signal.

**Remark.** Note that so far we study the potential sales, i.e., cart add-ins, as the dependent variable. When a customer purchases a lightning deal item that she has added to her cart, the purchase is counted as an actual sale and the incremental claim percentage that corresponds to her purchase remains in the inventory status bar. If a customer chooses not to check out the product, the item will be dropped from her cart and this dropped item will not translate into an actual sale. Thus, there is a difference between cart add-ins and actual sales. To show the robustness of our finding in actual sales, we also construct an approximate measure of the hourly sales as the difference between the hour-end and hour-beginning percentage claims. Please note that this measure is less accurate because it requires an implausible assumption that the exact same number of items are in
Table 5: Impact of Inventory Level

<table>
<thead>
<tr>
<th>Level of Scarcity</th>
<th>Claim &lt; 100%</th>
<th>Claim &lt; 85%</th>
<th>Claim &lt; 70%</th>
<th>Claim &lt; 55%</th>
<th>Claim &gt; 55%</th>
</tr>
</thead>
<tbody>
<tr>
<td>Claim</td>
<td>0.243***</td>
<td>0.225***</td>
<td>0.172***</td>
<td>0.150***</td>
<td>0.252***</td>
</tr>
<tr>
<td></td>
<td>(0.045)</td>
<td>(0.045)</td>
<td>(0.045)</td>
<td>(0.045)</td>
<td>(0.045)</td>
</tr>
<tr>
<td>Time control</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Characteristic control</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Observations</td>
<td>123,115</td>
<td>121,981</td>
<td>120,137</td>
<td>117,439</td>
<td>5,676</td>
</tr>
</tbody>
</table>

Note: This table reports the estimated coefficient values of equation (5) and standard errors clustered at the deal level with the time and deal characteristic controls over different levels of scarcity. The first four columns report the estimation based on subsets including deals with the maximum percentage-claim smaller than 100%, 85%, 70% and 55%, respectively. The last column reports the estimation based on a subset including deals with the maximum percentage-claim larger than 55%. The errors are corrected for heteroskedasticity. Significance at *p<0.1; **p<0.05; ***p<0.01.

the cart at the beginning and at the end of any time period that we study. Given that the average approximate sales is 2.4 units per hour, this measure becomes very sensitive to an imbalance between hour-beginning and hour-end in-cart units. Therefore, our main result does not include this measure. We repeat the fixed effect analysis and the interacted fixed effect analysis using this approximate measure, and we find that the signs of the parameters completely align with those for the cart add-ins. Nevertheless, if the conversion rate is constant across deals, our main finding with respect to the cart add-ins shows the ability of the inventory information to catch customers’ attention and translate more of these add-ins, these potential sales, into actual sales.

7.2.3. The impact of low inventory availability. To understand how consumer learning is impacted by the level of product availability, which captures the scarcity effect, we re-run the fixed-effect model across subsets of the panel data with various levels of scarcity. We measure the level of scarcity of each deal by its final claim. When the final percentage of products claimed is small, there is plenty of leftover inventory during the selling period and thus customers do not face a high out-of-stock risk. The dominant driver in customer learning is herding. In contrast, when the final claimed percentage is large, customers might rush their purchasing decisions because they might be afraid of stockout. The scarcity effect suggests that the momentum of customer learning increases as the percentage claimed increases.

Table 5 presents the estimated learning effects across subsets of the data with different scarcity levels. The results suggest a significantly increasing trend in the learning coefficient: the intensity of learning increases with the scarcity level. This empirical evidence is consistent with the scarcity effect. For deals with enough inventory, the learning momentum is likely to remain constant over time, while, for deals with low inventory, the learning momentum is intensified by the scarcity effect.
8. Conclusions
The extant literature in operations management has studied how customers’ strategic reactions to inventory information may impact firms’ inventory management, assortment, sourcing and information sharing decisions. Our paper provides empirical evidence that confirms the necessity to explore the strategic relations between customers and inventory information. In particular, our results show that consumers strategically learn from inventory information. There are two learning mechanisms. First, inventory information reflects the collective peer purchasing decisions, and consumers can learn about the deal quality from such information—creating an observational learning effect. Second, a low inventory availability creates an out-of-stock pressure among buyers, and thus prompts a resulting urgency to make an immediate purchase—leading to a scarcity effect. Our findings show evidence of both drivers.

Identifying the social learning effect has been recognized as a difficult task due to endogeneity concerns. To tackle this challenging problem, we design and run a randomized field experiment on Amazon by generating exogenous shocks to inventory information, which establishes the internal validity of our findings. We also employ a fixed-effect model analysis on our rich observational data to establish the external validity of our findings and study the underlying mechanisms of consumer learning. We find that consumers not only learn from the presence of herding, but also rationally moderate their inferences from the various reasons that give rise to the herd. In particular, our results show that consumers primarily learn about the deal’s value rather than the product quality from availability information while making a lightning deal purchase. Moreover, our empirical evidence suggests that scarcity may also drive purchasing decisions, as the learning momentum amplifies when availability decreases.

Because a decreasing inventory availability not only sends a favorable signal for product quality but also may create a stockout pressure among customers, our results have a limitation. First, we show that more past sales attract more future customers, which confirms that they react to and learn from inventory information. But we are unable to perfectly disentangle the herding and scarcity mechanisms. Future research should explore the conditions under which the scarcity effect begins to influence consumers and try to estimate herding and scarcity separately.

In an era when the advances in information technology have allowed companies to share their information at a much lower cost, it is important to understand how information impacts consumer behavior and reshapes demand so that it aligns profitably with supply. Information has been used as an important operations lever by companies to balance supply and demand. Prior research has shown that this goal can be achieved, for example, by disclosing obscure pricing or inventory information to customers (Fay and Xie 2008, Jerath et al. 2010, and Cui and Shin 2016), by sharing committed product availability to customers (Su and Zhang 2009), or by enhancing trust
in the shared information (Özer et al. 2011 and Özer et al. 2016). Our paper demonstrates that disclosing firms’ inventory availability information can also lead to customer learning and cause an acceleration in sales, and we hope that exploring such a relationship may spark future empirical research in the interaction between strategic customer learning and operations decisions.

Further, we are the first to collect data from a new and rich source: the Amazon lightning deal platform, which allows us to study consumer behavior in retail operations. We also point out a plausible and convenient way to run field experiments on Amazon’s flash sales platform. By showcasing this data source and the experimentation technique, we hope that our paper serves as a stepping stone for future research to explore new issues in consumer purchasing behavior.

References


Li, Xitong, Lynn Wu. 2014. Herding and social media word-of-mouth: Evidence from groupon. *Available at SSRN 2264411*.


Appendices

Figure 2  An Example of Percentage Claim Information over Time

Figure 3  Field Experiment Design Illustration
Figure 4  Field Experiment Mechanism Illustration

Click the “Add to Cart” button one time from one account; claim increases by 1%

The deal is added to the cart and can be found in the Shopping Cart

Click the “Add to Cart” button from 10 accounts within 2 minutes
## Table 6  Field Experiment Robustness Analysis

<table>
<thead>
<tr>
<th>Specification</th>
<th>Dependent variable: Cart add-ins per minute</th>
<th>Pre-50-min Comparison</th>
<th>Post-50-min Comparison</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>I. Treated</td>
<td>II. Control</td>
<td>III. DiD</td>
</tr>
<tr>
<td>Treated Period</td>
<td>0.045***</td>
<td>−0.012</td>
<td>−0.012</td>
</tr>
<tr>
<td></td>
<td>(0.017)</td>
<td>(0.017)</td>
<td>(0.017)</td>
</tr>
<tr>
<td>Treated Deal</td>
<td>−0.017</td>
<td>−0.012</td>
<td>−0.012</td>
</tr>
<tr>
<td></td>
<td>(0.018)</td>
<td>(0.018)</td>
<td></td>
</tr>
<tr>
<td>Treated Hour × Treated Deal</td>
<td>0.057**</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.026)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Saturday</td>
<td>0.004</td>
<td>0.082**</td>
<td>0.042*</td>
</tr>
<tr>
<td></td>
<td>(0.030)</td>
<td>(0.034)</td>
<td>(0.023)</td>
</tr>
<tr>
<td>Tuesday</td>
<td>0.040</td>
<td>0.040</td>
<td>0.033</td>
</tr>
<tr>
<td></td>
<td>(0.033)</td>
<td>(0.035)</td>
<td>(0.025)</td>
</tr>
<tr>
<td>Hour</td>
<td>−0.019**</td>
<td>0.026***</td>
<td>0.007</td>
</tr>
<tr>
<td></td>
<td>(0.008)</td>
<td>(0.008)</td>
<td>(0.006)</td>
</tr>
<tr>
<td>No. of Reviews</td>
<td>0.00003</td>
<td>0.0003***</td>
<td>0.0001***</td>
</tr>
<tr>
<td></td>
<td>(0.00003)</td>
<td>(0.0001)</td>
<td>(0.00003)</td>
</tr>
<tr>
<td>Review Rating</td>
<td>0.018**</td>
<td>0.007</td>
<td>0.015***</td>
</tr>
<tr>
<td></td>
<td>(0.007)</td>
<td>(0.007)</td>
<td>(0.005)</td>
</tr>
<tr>
<td>List discount</td>
<td>−0.0002</td>
<td>−0.0005</td>
<td>−0.001</td>
</tr>
<tr>
<td></td>
<td>(0.001)</td>
<td>(0.0004)</td>
<td>(0.0003)</td>
</tr>
<tr>
<td>Constant</td>
<td>0.379***</td>
<td>−0.362***</td>
<td>−0.008</td>
</tr>
<tr>
<td></td>
<td>(0.113)</td>
<td>(0.121)</td>
<td>(0.086)</td>
</tr>
<tr>
<td>Observations</td>
<td>362</td>
<td>528</td>
<td>890</td>
</tr>
</tbody>
</table>

Note: This table reports the estimated coefficient values of equations (1) and (2) with the time controls and deal characteristic controls. Column I reports the pre-post comparison relative to the pre-50-min period for the treated deals of equation (1); Column II reports the pre-post comparison relative to the pre-50-min period for the control group of equation (1); Column IV reports the pre-post comparison relative to the post-50-min period for the treated deals of equation (1); Column V reports the pre-post comparison relative to the post-50-min period for the control group of equation (1); Columns III and VI report the estimation results from the difference-in-difference approach in equation (2). The errors are corrected for heteroskedasticity. Significance at *p<0.1; **p<0.05; ***p<0.01.